

Urban transportation

* Transport economics: partial equilibrium, choice of commuting options.

* Urban wage: Commuting mode choice from the urban economics perspective.

→ full eq; but not as well suited for the data.

Commuting options: walk $w(L - b_w r) = z + sR(r)$

bus $w(L - X - b_B r) = z + sR(r)$

← drive $w(L - b_D r) - c = z + sR(r)$

Time cost important for high w

Pecuniary cost important for low w .

$b_w > b_B > b_D$ Time to travel \perp unit of distance.

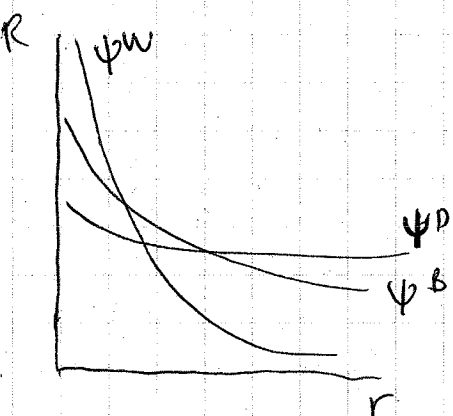
X → Fixed time cost for bus

c → Fixed pecuniary cost of driving.

Slopes of bid-rents

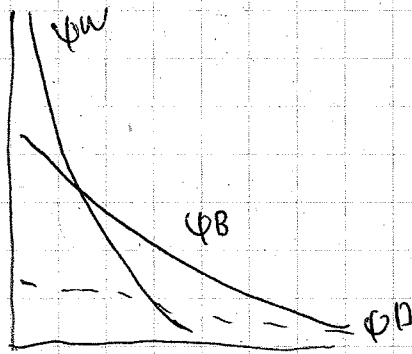
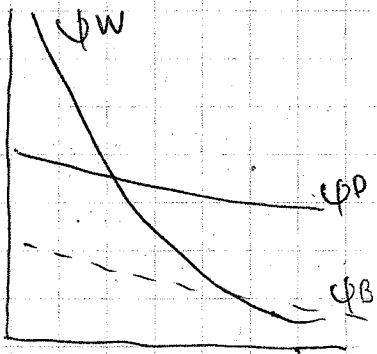
$$-\frac{wb_w}{s}, \quad -\frac{wb_B}{s}, \quad -\frac{wb_D}{s}$$

→ seems to match configuration of big cities...



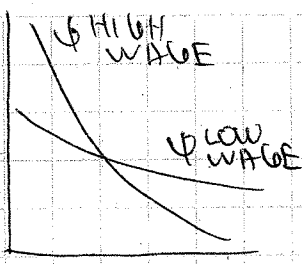
High wage

Low wage



* Why do the poor live in cities? should be outbid by the rich.

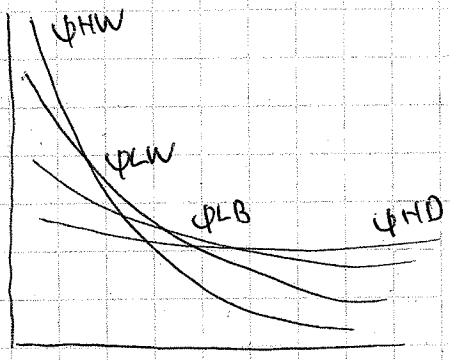
→ Coles paper Δ



if $\frac{d \ln s}{d \ln I} < 1$

↳ should hold within each commuting mode.

Combining the two



↳ Shows forces at play, but does not suit data well

Modelling individual commuting mode choice

- Mc Fadden (1974)

Random utility set-up

$$U(s, x) = v(s, f(x)) + \epsilon(s, x)$$

↳ indirect utility

Indiv choic
choice choic

deterministic

iid draw

↳ Prob of utility greater than other choic

$$\Pr [x_i | s, B] = \frac{1}{\prod_{i \neq j} \Pr [U(s, x_i) > U(s, x_j)]}$$

$$= \Pr [U(s, x_i) \geq \max_{j \neq i} U(s, x_j)]$$

Need to choose distribution with closed form for max.

~~Pr [E < Y]~~ $F(y) = \Pr [E < Y] = e^{-e^{-\frac{(y-M)}{b}}}$ Extreme value type I (Gumbel)

mode = M
mean = M + 0.58
sol = $\pi / \sqrt{6}$

median = $M - \ln(-\ln(2))$ } $b = 1$

$$\Pr [\text{choose 1}] = \Pr [v + \epsilon_1 \geq \max_{j > 1} [v_j + \epsilon_j]]$$

Distribution of max for EV type I: $\Pr [E_2 \leq v - v_2] \Pr [E_3 \leq v - v_3]$

$$\Pr [\max_{j > 1} [v_j + \epsilon_j] \leq v] = \prod_{j=2}^J e^{-e^{-(v-v_j)}}$$

$$= e^{-\sum_j e^{-v_j}} = e^{-e^{-v} \sum_j e^{v_j}}$$

$$\Pr[\max_{j>1} [v_j + \epsilon_j] \leq u] = e^{-e^{-v} \sum_{j>1} v_j}$$

$$\text{So } \Pr[v_1 + \epsilon_1 \geq \max_{j>1} [v_j + \epsilon_j]] = e^{-e^{-(v_1 + \epsilon_1)} \sum_{j>1} v_j}$$

Unconditional probability of choosing 1

$$= \int_{-\infty}^{\infty} f(\epsilon_1) \Pr(v_1 + \epsilon_1 \geq \max_{j>1} [v_j + \epsilon_j]) d\epsilon_1$$

$$= \frac{e^{-v_1} (e^{v_2} + e^{v_3} + \dots + e^{v_K}) + 1}{e^{v_1} (e^{v_2} + e^{v_3} + \dots + e^{v_K}) + 1} = \frac{e^{v_1}}{\sum_{j=1}^K e^{v_j}}$$

switch in notation

$$P_{ij} = \frac{\exp(\phi_i^T s_i + \theta^T c_{ij} s_j)}{\sum_K \exp(\phi_j^T s_j + \theta^T c_{kj} s_j)}$$

choice ↙ ↘ chars

Estimate by ML

$$\sum_i \phi_i = \prod_j [\prod_i [P_{ij}]^{P_{ij}}] = 1 \text{ if } j \text{ chooses } i$$

0 or w

j → "what units"? Households, individuals? ↗ IIA household allocation Δ
 i → IIA, criticism.
 IIA may not hold given the information we have about choices.

Identification: Not all ϕ, θ identified, only can identify them relative to a reference choice.

$$P_{ij} = \frac{\exp[(\phi_i - \phi_1)^T s_i + \theta^T (c_{ij} - c_{1j}) s_j]}{1 + \sum_{k>1} \exp[(\phi_k - \phi_1)^T s_i + \theta^T (c_{kj} - c_{1j}) s_j]}$$

What to do to get at IIA?
 → Normal disturbances
probit random coefficients model

Probit random coefficients model

* Correlation in errors and coefficients across choices.

Suppose $U_i = Z_i \beta + \eta_i$ for choice i \rightarrow can be relaxed.

$$\eta_i = N(0, \sigma^2_i), \quad \eta_i \perp Z_i, \eta_j, \beta$$

$$\beta \sim N(\bar{\beta}, \Sigma_\beta)$$

$$U_i = Z_i \bar{\beta} + Z_i (\beta - \bar{\beta}) + \eta_i$$

Suppose there are three choices

If $U_1 \geq U_2$

$$U_1 - U_2 = (Z_1 - Z_2) \bar{\beta} + (Z_1 + Z_2) (\beta - \bar{\beta}) + \eta_1 - \eta_2 \geq 0$$

$$\text{Var}(U_1 - U_2) = (Z_1 - Z_2) \Sigma_\beta (Z_1 - Z_2)^T + \sigma_1^2 + \sigma_2^2$$

$$\text{Cov}(U_1 - U_2, U_1 - U_3) = (Z_1 - Z_2) \Sigma_\beta (Z_1 - Z_3)^T + \sigma_1^2$$

$$\rho = \text{corr}(U_1 - U_2, U_1 - U_3) = \frac{\text{Cov}(U_1 - U_2, U_1 - U_3)}{\sqrt{\text{Var}(U_1 - U_2) \text{Var}(U_1 - U_3)}}$$

want to know $\Pr\{I_1 | I_1, 2, 3\}$

$$= \Pr\{U_1 - U_2 \geq 0 \text{ and } U_2 - U_3 \geq 0\}$$

$$= \Pr\left\{ \begin{aligned} (Z_1 - Z_2) \bar{\beta} + \sqrt{\text{Var}(U_1 - U_2)} \epsilon_1 &\geq 0 \\ \text{and } (Z_2 - Z_3) \bar{\beta} + \sqrt{\text{Var}(U_2 - U_3)} \epsilon_2 &\geq 0 \end{aligned} \right\}$$

ϵ_1, ϵ_2 joint standard normal \uparrow

\rightarrow specify correlation later.

$$\Pr(\text{choose } 1) = \int \int f(t_1, t_2) dt_2 dt_1$$

Limits

$$= \int_{-\infty}^{\frac{(Z_1 - Z_2) \bar{\beta}}{\sqrt{\text{Var}(U_1 - U_2)}}} \int_{-\infty}^{\frac{(Z_2 - Z_3) \bar{\beta}}{\sqrt{\text{Var}(U_2 - U_3)}}} f(t_1, t_2) dt_2 dt_1$$

$$J(\bar{\beta}, \Sigma_\beta, \sigma_1^2, \sigma_2^2, \sigma_3^2) = \prod_j \pi_j ([\Pr \text{ choose } i]_j)^{D_{ij}}$$

Small / Winston / Van

"Revealed preference"

Not enough variation

"Stated preference"

F or X
cost on X
relative time
time of day →

relative 80-50 pctile
in time.

Variables:

c_{it} toll difference
 R_{it} unreliability difference
 T_{it} time difference

Values of travel time:

$$\frac{\partial V/\partial T}{\partial V/\partial C} = \frac{\partial V/\partial T}{\lambda} = \frac{\text{utils/time}}{\text{utils/\$}} = \$/\text{time}$$

$$U_1 = v_1 + \epsilon_1$$

$$U_2 = v_2 + \epsilon_2$$

$$\Pr(1 | 1, 2) = \frac{e^{v_1}}{e^{v_1} + e^{v_2}} = \frac{e^{v_1 - v_2}}{1 + e^{v_1 - v_2}}$$

$$\Pr(2 | 1, 2) = \frac{1}{1 + e^{v_1 - v_2}}$$

$$v_1 - v_2 = u_i - n_i$$

$$\Pr(U_1 > U_2) = \Pr$$

$n_i \sim \text{PEV}$
 $Dn_i \sim \text{probit}$

$$U_{it} = \theta_i + \beta_i X_{it}$$

choice chars

$$VOT_i = \frac{\partial U_{it} / \partial T_{it}}{\partial U_{it} / \partial C_{it}}$$

$$VOR_i = \frac{\partial U_{it} / \partial R_{it}}{\partial U_{it} / \partial C_{it}}$$

Heterogeneity:

$$\theta_i = \bar{\theta} + \phi w_i + \epsilon_i \sim N(\bar{\theta}, \sigma_\epsilon^2)$$

$$\beta_i = \bar{\beta} + \tau z_i + \eta_i \sim N(0, \sigma_\eta^2)$$

Some more assumptions on error structure, final system

Estimation

Example for two choices: logit

\downarrow x_{priest} \downarrow freeway
 $1 = X$ $2 = F$

$$U_1 = X_1 \beta + \eta_1$$

$$U_2 = X_2 \beta + \eta_2$$

$$\rightarrow Pr[X] = Pr[(X_1 - X_2) \beta > \eta_2 - \eta_1]$$

$$= \frac{e^{(X_1 - X_2) \beta}}{1 + e^{(X_1 - X_2) \beta}}$$

$$f = \prod_j \left[\frac{e^{(X_{1j} - X_{2j}) \beta}}{1 + e^{(X_{1j} - X_{2j}) \beta}} \right]^{D_j} \left[\frac{1}{1 + e^{(X_{1j} - X_{2j}) \beta}} \right]^{1 - D_j}$$

Their model

$$f = \prod_j \int_{\Theta} \prod_{i \in \text{obs}} Pr(X_i | \phi, \theta)^{D_{ij}} Pr(F_i | \phi, \theta)^{1 - D_{ij}} f(\theta) d\theta$$

ϕ
 Individuals
 \downarrow
 obs
 within
 individual

choices independent
 conditional on
 the random effects
 from Θ .

$$Pr(X_i | \phi, \theta) = F[\theta^{BR} + \phi^{BR} w_i + \beta_i^{BR} X_i^{BR} + v_i^{BR}]$$

or
 $F[BS]$

c does not enter integral, only one obs, not correlated.

Algorithm

1. θ^0
2. Simulate L , get θ^+ by gradient based method
3. Stop when $|\theta^{k+1} - \theta^k| < \text{stopping crit.}$

Policy paper \rightarrow NOV simulation